



SOURCE CONTRACT CONTRACTOR CONTRACTOR CONTRACTOR

	$\overline{}$
_	<b>~</b> `
	J
✓-	₹,

SECURITY CLASSIFICATION OF THIS PAGE	<u></u>	Σ.						
	REPORT DOCUME	ENTATION PAGE	<b>.</b>					
1. REPORT SECURITY CLASSIFICATION Unclassified		1b. RESTRICTIVE MARKINGS						
2& SECURITY CLASSIFICATION AUTHORITY N/A		3. DISTRIBUTION/AVAILABILITY OF REPORT APPROVED FOR PUBLIC RELEASE: DISTRIBUTION UNLIMITED						
26. DECLASSIFICATION/DOWNGRADING SCHEDULE N/A		DISTRIBUTION UNDITED						
4. PERFORMING ORGANIZATION REPORT NUMBERIS)		5. MONITORING ORGANIZATION REPORT NUMBER(S)  AFOSR-TR- 86-0377						
64 NAME OF PERFORMING ORGANIZATION UNIVERSITY OF MASSACHUSETTS	6b. OFFICE SYMBOL (If applicable)	7. NAME OF MONITORING ORGANIZATION AFOSR/NM						
Oc. AODRESS (City. State and ZIP Code) Department of Mathematics & Statistics Lederle Graduate Towers Amherst, MA 01003		7b. ADDRESS (City, State and ZIP Code:  Building 410  Bolling Air Force Base, DC 20332-6448						
84. NAME OF FUNDING/SPONSORING ORGANIZATION AFOSR	8b. OFFICE SYMBOL (If applicable) NM	9. PROCUREMENT INSTRUMENT IDENTIFICATION NUMBER AFOSR-82-0167						
Bc. ADDRESS (City, State and ZIP Code)		10. SOURCE OF FUR	NDING NOS.					
Building 410 Bolling Air Force Base, DC 20332-6448		PROGRAM ELEMENT NO.	PROJECT NO.			WORK UNIT		
11. TITLE (Include Security Classification) Approximate Counting: A Marti	6.1102F	2304						
12. PERSONAL AUTHORIS) W. A. Rosenkrantz	•			_				
	5/35/05 5/34/0			14. DATE OF REPORT (Yr., Mo., Day) 15. PAGE COUNT 17, 1986 13				
16. SUPPLEMENTARY NOTATION								
17. COSATI CODES	18 SUBJECT TERMS	18. SUBJECT TERMS (Continue on reverse if necessary and identify by block number)						
FIELD GROUP SUB. GR.	Approximate Counting, Martingales, Exponential Back Off Protocol							

19. ABSTRACT (Continue on reverse if necessary and identify by block number

Approximate counting is a probabilistic algorithm for keeping track of large numbers of events by means of counters of limited range. In this paper we present an analysis of this algorithm using the elementary theory of martingales. The methods are also applicable to the analysis of the counter which occurs in the exponential back off protocol.

## OTIC FILE COPY

SELECTE JUL 2 4 1986

20 DISTRIBUTION/AVAILABILITY OF ABSTRACT	21. ABSTRACT SECURITY CLASSIFICATION			
UNCLASSIFIED/UNLIMITED 🛭 SAME AS APT 🗍 OTIC USEAS 🗎	UNCLASSIFIED			
27. NAME OF RESPONSIBLE INDIVIDUAL	226 TELEPHONE NUMBER	22c OFFICE SYMBOL		
BRIAN W. WOODRUFF, MAJ.	(202) 767-5027	AFOSR/NM		

DD FORM 1473, 83 APR

DITION OF LIAN 73 IS OBSOLETE

**NUCLY221-IFD** 

86

# APPROXIMATE COUNTING : A MARTINGALE APPROACH

Walter A. ROSENKRANTZ (1)
INRIA

Domaine de Voluceau BP 105 - Rocquencourt 78153 Le Chesnay Cedex - France MOTICE OF TRANSMITTAL TO DTIC
This technical report has been reviewed and is
"Approved for public release IAW AFR 190-12.
"ATTHEM J. KENTER
Chief, Technical Information Division

### I - INTRODUCTION

In the paper "Counting Large Numbers of Events in Small registers", Morris [Mo 1978] proposed a probabilistic algorithm for keeping track of a large number of events M with an n bit binary counter where typically M  $>> 2^n-1$  = largest integer that can be represented by the counter. This method of counting has been dubbed "Approximate counting" by Flajolet [Fl 1985] who reformulated the problem in terms of a discrete time Markov chain b(t) with state space

$$I^+ = \{0, 1, 2, \ldots\}$$

and transition function given by

(1.1) 
$$\begin{cases} P(b(t+1) = i+1 \mid b(t) = i) = 2^{-i}, & i \ge 0 \\ P(b(t+1) = i \mid b(t) = i) = 1-2^{-i}. \end{cases}$$

The process b(t) arises naturally when one "counts" an event with probability  $2^{-b(t)}$  and does not record it with probability  $1-2^{-b(t)}$ , where b(t) = current count in the register. How well does b(t) track t?

Morris [Mo 1978] has noted that

Approved for public release;

<sup>(1)</sup> Research supported by AFOSR Grant #82-0167

(1.2) 
$$\begin{cases} a) & E(2^{b(t)}) = t+2 \\ b) & \sigma^2(2^{b(t)}) = t(t+1)/2 \end{cases}$$

and Flajolet [Fl 1985] has given a proof of (1.2).

Thus  $2^{b(t)}$ -2 is an unbiased estimator of t. In addition Flajolet has shown, Theorem 1 of [Fl 1985], that

(1.3) 
$$a_1(t) \le E(b(t)) - \log_2 t \le a_2(t)$$

where  $a_i(t)$  are small and bounded as  $t \rightarrow \infty$ .

The proof of (1.3) given in [Fl 1985] is not simple since it uses Mellin transforms and other refined techniques from the theory of functions of a complex variable. It is the purpose of this paper to derive (1.2) and (1.3) by means of a more elementary method, at least more elementary to probabilists !, using only the simplest ideas of the theory of martingales. See chapter VI of [Ka-Ta 1975] for a more comprehensive account. More precisely, we shall prove that (1.3) holds with

$$a_1(t) = \log_2 (2t^{-1} + \log_2)$$

and

$$a_2(t) = \log_2 (1 + 2t^{-1}).$$

In addition we derive a new recursive formula for  $E(2^{lb(t)})$ ,  $l \ge 1$ .

Interesting/pg enough the same process b(t), but with b(1) = 0, occurs in a recent paper by [GGMM 1985] where b(t) represents the back off counter occuring in the exponential back off protocol (EBO) when the channel is always jammed; in fact, this is what first stimulated my interest in this problem - see [Ro 1984]. The problem here is to show that

$$\lim\sup_{t\to\infty} t \, \mathbb{E}(2^{-b(t)}) < v.$$

In fact the following result was proved in [GGMM 1985]

Availability Codes

Avail and for Special

A-1

(1.4) 
$$1/8 \le t E(2^{-b(t)}) \le 9$$
, for all  $t \ge 1$ .

By a simple martingale argument we are able to obtain the sharper lower bound

$$(1.5) 1 \le t E(2^{-b(t)})$$

and with a little more effort we are able to show

(1.6) 
$$t E(2^{-b(t)}) \le 4.1, t \ge 1.$$

In addition [GGMM 1985] conjectured that

(1.7) 
$$E(2^{-b(t)}) \sim c t^{-1}$$
 where  $c = 1/\log 2$ .

In fact Flajolet has shown [Fl 1986] that (1.7) must be modified tb take into account bounded fluctuations  $\omega(t)$  of small amplitude i.e. he shows that

(1.8) 
$$E(2^{-b(t)}) = c t^{-1} + \omega(t)$$
.

The proof of (1.8) is, as is to be expected, quite delicate. Using only very simple tools we are able to prove the following weak form of conjecture (1.7)

(1.9) 
$$\begin{cases} a) & \limsup \Sigma & E(2^{-b(s)})/\log t \le c \\ & t \to \infty & s=1 \end{cases}$$

$$\begin{cases} t-1 & E(2^{-b(s)})/\log t \le c \\ & t \to \infty & s=1 \end{cases}$$

These results are obtained in part 3.

Acknowledgments: I wish to thank my colleagues F. Baccelli, G. Fayolle, P. Flajolet and P. Robert here at INRIA for several helpful discussions. In particular the proof of the lower bound (2.9) is due to P. Robert.

#### II - A METHOD FOR CONSTRUCTING MARTINGALES ASSOCIATED WITH A MARKOV CHAIN

The following discussion assumes the reader is familiar with chapter VI of Karlin-Taylor. Let y(t) be a Markov chain with countable state space  $I^+ = \{0, 1, 2, ...\}$  and transition matrix P = (P(i,j)). Let  $\mathscr{F}(t) = \mathscr{B}(y(u), 0 \le u \le t)$  denote the smallest  $\sigma$ -field generated by the random variables y(u),  $0 \le u \le t$ . Intuitively  $\mathscr{F}(t)$  is the past up until time t. We recall the

In order to construct martingales associated with the Markov chain y(t) we make use of the operators P and A acting on the function space

$$D = \{f \mid f: I^+ + R, \sum_{j=0}^{\infty} P(i,j) \mid f(j) \mid < \infty \}$$

Definition : Let f € D

(i) 
$$Pf(i) \Delta \sum_{j=0}^{\infty} P(i,j) f(j)$$

(2.2) (ii) Af(i) 
$$\Delta$$
 Pf(i) - f(i)

(iii) 
$$P|f|(i) \triangleq \sum_{j=0}^{\infty} P(i,j)|f|(j).$$

Note that  $E\{f(y(t+1)) - f(y(t)) | \mathcal{F}(t)\} = E\{f(y(t+1)) - f(y(t)) | y(t)\} = Af(y(t))$  and using this fact it is easy to establish the following

Lemma : Suppose  $P[f](i) < \infty$  and set

(2.3) 
$$x(t) = f(y(t)) - \sum_{s=0}^{t-1} Af(y(s)), t \ge 1 \text{ and } x(0) = f(y(0)).$$

Then  $\{x(t), \widehat{\mathcal{J}}(t)\}\$  is a martingale.

For future reference we note the following special case. Suppose Af(i) = c all  $i \ge 0$ . Then f(y(t)) - ct is a martingale.

Lemma (2.3) yields a novel and computationally simple proof of Flajolet's proposition 0 as well as a recursive formula for the higher moments of  $2^{b(t)}$ .

Proposition:

(2.4)

Let 
$$b(0) = 1$$
. Then  $E(2^{b(t)}) = t+2$  and  $\sigma^2(2^{b(t)}) = t(t+1)/2$ .

Proof. We begin by noting that

(2.5) 
$$Af(j) = 2^{-j} [f(j+1) - f(j)].$$

Let  $f_{\ell}(j) = 2^{\ell j}$  where  $\ell = 0, 1, 2, \ldots$  An easy calculation shows that

(2.6) 
$$Af_{\ell}(j) = (2^{\ell-1}) f_{\ell-1}(j).$$

In particular  $Af_1(j) \equiv 1$  and therefore, by lemma (2.3), we have

$$f_i(b(t)) - \sum_{s=0}^{\ell-1} 1 = 2^{b(t)} - t$$

is a martingale. Consequently

$$E(f_1(b(t)) - t) = f_1(b(0)) = f_1(1) = 2$$
 i.e.

(2.7) 
$$E(f_1(b(t)) = E(2^{b(t)}) = t+2.$$

Lemma (2.3) and formula (2.6) can now be used to derive a recursive formula for the moments  $E(2^{lb(t)})$ ,  $l \ge 2$ . To illustrate these ideas consider the case l = 2. Since  $Af_2(j) = 3f_1(j)$  lemma (2.3) and (2.7) yield

$$E(f_{2}(b(t)) - \sum_{s=0}^{t+1} 3f_{1}(b(s)) = f_{2}(b(0)) = 4$$

$$E(2^{2b(t)}) = E(f_2(b(t)) = 4 + 3 \sum_{s=0}^{t-1} E(f_1(b(s)) = 4 + 3 \sum_{s=0}^{t-1} (s+2) = s = 0$$

$$= 4 + 3 \cdot \frac{(t-1)t}{2} + 6 \cdot t = \frac{3}{2} t(t+3) + 4.$$

Now

$$\sigma^2(2^{b(t)}) = E(2^{2b(t)}) - E(2^{b(t)})^2 = \frac{3}{2} t(t+3) + 4 - (t+2)^2 = \frac{t(t+1)}{2}.$$

Using (2.3) we can give a recurrence formula for the higher moments of  $2^{b(t)}$ . More precisely let  $M_{\ell}(t) = E(2^{\ell b(t)})$ ,  $\ell = 0, 1, \ldots$ . Then lemma (2.3) and (2.6) yield

(2.8) 
$$M_{\ell}(t) = 2^{\ell} + (2^{\ell-1}) \sum_{s=0}^{t-1} M_{\ell-1}(s).$$

Note. We believe formula (2.8) to be new.

The formula  $E(2^{b(t)})$  = t+2 together with Jensen's inequality applied to the concave function  $\phi(t)$  =  $\log_2 t$  yield

$$E(b(t)) \le \log_2(t+2) = \log_2 t + \log(1+2t^{-1}).$$

We now proceed to derive a lower bound for  $\mathrm{E}(2^{\mathrm{b}(t)})$  using an argument suggested to the author by P. Robert (INRIA).

An easy calcullation shows that

$$E(b(t+1) - b(t)) = E(2^{-b(t)}) \ge 2^{-E(b(t))}$$

and therefore

$$2^{E(b(t))}$$
 .  $E(b(t+1) - b(t)) \ge 1$ .

Thus

Set  $\phi(s) = E(b(s))$ ; so  $\phi(0) = 1$ . The preceeding expression can be rewritten as

$$\begin{array}{l} t^{-1} \\ \Sigma \\ z^{\phi(s)} \end{array} (\phi(s+1) - \phi(s)) \ge t.$$
s=0

On the other hand

$$t \le \sum_{s=0}^{t-1} 2^{\phi(s)} \left[ \phi(s+1) - \phi(s) \right] \le \int_{1}^{\phi(t)} 2^{s} ds = e^{(\log 2)\phi(t)} - e^{\log 2} / \log 2 \text{ i.e.}$$

$$2^{\phi(t)} \ge 2 + (\log 2).t$$

and this implies

$$E(b(t)) = \phi(t) \ge \log_2 (t[\log_2 + 2t^{-1}]) = \log_2 t + \log_2 [\log_2 + 2t^{-1}].$$

Summing up then we're shown that

$$\log_2 t + \log_2 (2t^{-1} + \log_2) \le E(b(t)) \le \log_2 t + \log_2 (1 + 2t^{-1}).$$

Clearly this implies  $\lim_{t\to\infty} \frac{E(b(t))}{\log_2 t} = 1$  and also that

(2.10) 
$$\log_2(\log 2) \le E(b(t)) - \log_2 t \le 1/t$$
.

### III - AN IMPROVED ESTIMATE FOR E(2-b(t))

In this part of the paper we assume b(1) = 0 and repeating the same argument of proposition (2.4) we see that

$$f_1(b(t)) - \sum_{s=1}^{t-1} Af_1(b(s))$$

is a martingale with

$$E(f_1(b(t)) - \sum_{s=1}^{t-1} Af_1(b(s)) = f_1(b(1)) = 1.$$

Consequently  $E(2^{b(t)}) = t$ . We now apply Jensen's inequality with  $p(x) = x^{-1}$  to deduce the much sharper lower bound

(3.1) 
$$E(2^{-b(t)}) \ge 1/t$$
.

To obtain the upper bound

(3.2) 
$$E(2^{-b(t)}) \le 4.1/t$$
.

We derive an exponential bound on the first passage time  $\tau_y$  = inf{t : b(t)  $\geq$ y}. This leads to a sharper estimate than the one obtained by [GGMM 1985] who used Chebychev's inequality; otherwise the two proofs are the same.

Lemma: For any 
$$\lambda \le \lambda(y) = -\log(1 - 2^{-y-1})$$

(3.3)
$$P(\tau_{y} \ge t) \le \exp(-\lambda t + \lambda) \prod_{i=1}^{y-1} (1 - (1 - \exp(-\lambda))2^{i})^{-1}.$$

<u>Proof.</u> Let  $y_i$  be the random variable with geometric distribution given by  $P(y_i=j)=2^{-i}(1-2^{-i})^{j-1}$ ,  $j\geq 1$ . Thus its moment generating function

$$E(e^{\lambda y}i) = (1 - (1 - exp(-\lambda))2^{i})^{-1}.$$

Clearly  $\tau_1 = 1$  and for y > 1 we have

Since the  $y_i$ , 1  $\leq$  i  $\leq$  y-1, are mutually independent we see at once that

$$M_{y}(\lambda) = E(\exp(\lambda \tau_{y})) = \exp(\lambda) \prod_{i=1}^{y-1} E(\exp(\lambda y_{i}))$$

$$= \exp(\lambda) \prod_{i=1}^{y-1} (1 - (1 - \exp(-\lambda))2^{i})^{-1}.$$

By Chebychev's inequality  $\exp(\lambda t)$   $P(\tau_y \ge t) \le M_y(\lambda)$  which is precisely lemma (3.3). Note that a sufficient condition for

$$M_{y}^{(\lambda)} < \infty \text{ is } \sup_{1 \le i \le y-1} 2^{i} (1 - \exp(-\lambda)) < 1.$$

In particular if we set  $\lambda = \lambda(y)$  where

(3.4) 
$$\lambda(y) = -\log(1 - 2^{-y})$$

then  $M_y(\lambda) < \infty$  for  $\lambda \le \lambda(y)$  and

(3.5) 
$$\log(1 - (1 - \exp(-\lambda))2^{i}) \ge - (\log 2)(1 - \exp(-\lambda))2^{i},$$

where we've used the inequality

$$log(1-x) \ge -(log2)x$$
 on  $0 \le x \le \frac{\pi}{2}$ .

Our next step is to set  $\lambda = \lambda(y)$  in lemma (3.3) and then take the log of both sides which leads to

$$\log P(\tau_{y} \ge t) \le -\lambda(y)(t-1) + (\log 2) \sum_{i=1}^{y-1} 2^{i-y}$$

$$\le -\lambda(y)t + (\log 2)(1 - 2^{-y}) + \lambda(y).$$

Consequently

$$P(\tau_y \ge t) \le c(y) \exp(-\lambda(y)t)$$

where

TO SERVICE STATES STATES OF THE SERVICES OF

$$c(y) = \exp{\lambda(y) + (\log 2)(1 - 2^{-y})} \le 2^{3/2}, y \ge 1.$$

We have thus derived the exponential bound

(3.6) 
$$P(\tau_y \ge t) \le 2^{3/2} \exp(-\lambda(y)t).$$

We now proceed to estimate  $E(2^{-b(t)})$ . Set  $m = \lfloor \log_2 t \rfloor$  so  $(\log_2 t)-1 < m \leq \log_2 t$ , consequently

$$E(2^{-b(t)}; b(t) \ge m) \le 2/t.$$

Here  $\mathsf{E}(\mathsf{f};\mathsf{A})$  means integrating  $\mathsf{f}$  over the subset  $\mathsf{A}.$  So the problem row is to show that

$$E(2^{-b(t)}; b(t) \le m-1) \le a/t$$

and to estimate the constant a. Now

(3.7) 
$$E(2^{-b(t)}; b(t) \le m-1) \le \sum_{j=0}^{m-1} 2^{-j} P(b(t) \le j).$$

But

$$P(b(t) \le j) = P(\tau_j \ge t) \le 2^{3/2} \exp(-\lambda(j)t)$$
  
-  $\lambda(j) = + \log(1 - 2^{-j}) \le -2^{-j}$ 

together imply

and

PROCESSES PROCESSES PROCESSES PROCESSES PROCESSES PROCESSES

$$P(b(t) \le j) \le 2^{3/2} \cdot exp(-2^{-j}t)$$
  
 $\le 2^{3/2} exp(-2^{m-j}).$ 

The right hand side of (3.7) is bounded by the sum

$$S_1 = \sum_{j=0}^{m-1} 2^{-j} \exp(-2^{m-j}).$$

Since -  $2^{m-j} \le - (m-j)2$  it follows that

$$S_1 \le 2^{3/2} \frac{m^{-1}}{\sum_{j=0}^{\infty}} 2^{-j} [\exp(-2)]^{m-j} = 2^{3/2} \exp(-2m) \frac{m^{-1}}{\sum_{j=0}^{\infty}} {\frac{\exp(2)}{2}}^{j} \le$$

$$\le (2^{3/2} / [(e^2/2)^{-1}]) 2^{-m} \le c/t$$

where

$$a = 2^{5/2} / [(e^2/2)-1] = 2.0993859.$$

Finally, putting these estimates together we obtain

(3.8) 
$$E(2^{-b(t)}) \le 4.0993859 / t.$$

We conclude this part of the paper by establishing a weak form of the conjecture (1.7). More precisely set  $m(t) = E(2^{-b(t)})$  and apply lemma (2.3) to the function f(j) = j. Since  $Af(j) = 2^{-j}$  we see at once that

$$E(b(t) - \sum_{s=1}^{t-1} 2^{-b(s)}) = 0$$

and therefore

COLORGE PRODUCTS CONTROL CONTROL OF

On the other hand as we've already seen  $E(2^{-b(t)}) \ge 1/t$  which implies

(3.10) 
$$\sum_{s=1}^{t-1} E(2^{-b(s)}) \ge \sum_{s=1}^{t-1} \frac{t-1}{s} \ge 1 + \int_{s}^{t-1} \frac{ds}{s} = 1 + \log(t-1).$$

Combining (3.9) and (3.10) yields

$$t=1$$

$$\sum_{\Sigma} E(2^{-b(s)})$$
(a)  $\limsup_{t \to \infty} \frac{s=1}{\log t} \le 1/\log 2$ 

(3.11) 
$$\begin{array}{ccc} t=1 & & \Sigma & E(2^{-b}(s)) \\ & \Sigma & E(2^{-b}(s)) & & \Sigma & E(2^{-b}(s)) \end{array}$$
 (b) 
$$\begin{array}{cccc} t=1 & & & \Sigma & E(2^{-b}(s)) & & \Sigma$$

Note that if in fact  $E(2^{-b(s)})$  c/s then  $\Sigma$   $E(2^{-b(s)})$  - c log t. This is s=1 why we call (3.11) the weak form of conjecture (1.7).

#### IV - AN ESTIMATE FOR THE TAIL OF THE DISTRIBUTION

Using the result that  $E(2^{2b(t)}) = \frac{4}{2} + (3/2) t(t+3)$  and Chebychev's inequality it is easy to see that

(4.1) 
$$P(b(t) \ge 2 \log_2 t + \delta) = O(2^{-2\delta} t^{-2}), \ \delta \ge 0.$$

In fact

$$P(b(t) \ge 2 \log_2 t + \delta) = P(2^{2b(t)} \ge 2^{2\delta}t^*) \le \frac{4 + (3/2)t(t+3)}{2^{2\delta} \cdot t^*} \le c_1 2^{-2\delta} \cdot t^{-2}$$

where  $c_1$  is independent of  $\delta$  and t. This is in fact a considerable strengthening of Flajolet's proposition 4 wherein he shows that

$$P(b(t) = 2\log_2 t + \delta) = 0(2^{-\delta}t^{-.99}),$$

uniformly in t and  $\delta \ge 0$ .

16

#### REFERENCES

- [1] [F1 1985] P. Flajolet: "Approximate counting: a detailed analysis". BIT 25 (1985), p. 113-134.
- [2] [F1 1986] P. Flajolet: Personal communication.

MODERAN CONTRACT MANAGEMENT CONTRACTOR SECRETARY

- [3] [GGMM 1985] J. Goodman, A. Greenberg, N. Madras, P. March: "On the stability of the ethernet". Preprint April 85.
- [4] [Ka-Ta 1975] S. Karlin, H. Taylor: "A first course in stochastic processes". 2nd ed. Academic Press, New York (1975).
- [5] [Mo 1978] R. Morris: "Counting large numbers of events in small registers". Comm. ACM, 21, (1978). pp. 840-842.
- [6] [Ro 1984] W.A. Rosenkrantz: "Some theorems on the instability of the exponential back-off protocol". 10th International Symposium Computer Performance, Paris, France (Dec. 1984), 199-205.

ALTA CONTRACTOR